

Restraint use as a quality indicator for the hospital setting: a secondary data analysis

Silvia Thomann^a, Sabine Hahn^a, Kai-Uwe Schmitt^{bc}, Isabelle Barbezat^c, Sandra Siegrist-Dreier^a, Dirk Richter^{ade}

^a Bern University of Applied Sciences, Department of Health Professions, Applied Research and Development in Nursing, Bern, Switzerland

^b Academic Practice Partnership of Bern University of Applied Sciences and University Hospital Bern (Insel Group), Bern, Switzerland

^c Bern University Hospital (Insel Group), Department of Nursing, Bern, Switzerland

^d Bern University Hospital for Mental Health, Centre for Psychiatric Rehabilitation, Bern, Switzerland

^e University of Bern, University Hospital for Psychiatry and Psychotherapy, Bern, Switzerland

Summary

INTRODUCTION: A reduction in restraint use is recommended for all health care settings. For this purpose, local or national quality measurement and improvement initiatives have been implemented in various countries, primarily in the mental health and long-term care settings. However, restraints are also frequently used in the somatic acute care hospital setting, and strong variations in their prevalence rates have been reported. Therefore, the aim of this study was to reanalyse existing data on restraint use in Swiss hospitals in order to assess the potential of restraint use as a national quality indicator for the hospital setting.

METHODS: Using a cross-sectional, multicentre design, data were collected between 2016 and 2018 as part of the ANQ's (Swiss National Association for Quality Development in Hospitals and Clinics) "falls and pressure ulcers" national prevalence measurement in acute care hospitals in Switzerland. The hospitals measured restraint use on a voluntary basis in addition to falls and pressure ulcers. All medical specialities and patients aged 18 and over who gave their informed consent were included in the measurement. Descriptive and multilevel regression analyses were performed using institutional, ward and patient-level data relating to restraint use.

RESULTS: The sample consisted of 18,938 inpatients from 55 hospitals. The 30-day prevalence rate of patients with at least one restraint was 10.2% (n = 1933). The risk-adjusted hospital comparison revealed that hospitals in Switzerland differ significantly in their restraint use, even after adjusting for patient characteristics. In total, 10 hospitals used restraints significantly less and 12 used them significantly more than the national average.

CONCLUSION: Restraint use varies significantly between Swiss hospitals: 40% of all hospitals used restraints either significantly more or significantly less often than the average. In comparison to the other quality indicators, this is a very high value, indicating potential for improvements in the quality of care. Since restraint use is associated not only with quality of care, but also with human rights, these

large differences seem questionable from a professional, ethical and legal point of view. Clearer and binding regulations in combination with monitoring and benchmarking of restraint use in hospitals, such as with a national quality indicator, seem necessary. These would help to ensure that restraint use is in alignment with professional values, as well as ethical and legal requirements.

Introduction

Restraints have been used in health care settings for centuries. In mental health care, there is an increasing awareness of the negative consequences of restraints, and therefore, restraint use is more and more regulated. For example, it has been an important quality indicator for inpatient psychiatry for many years [1–3]. This restrictive practice is increasingly also viewed critically in nursing home settings. In Australia, for example, stricter regulations regarding restraint use in residential care settings were introduced in 2019 [4]. In Switzerland, interventions restricting movement in nursing homes have been monitored at the national level since 2019 [5]. However, most countries lack clear regulations for somatic acute care hospitals (subsequently referred to as "hospitals"), and often only recommendations exist. In Canada, for example, recommendations from the Canadian Agency for Drugs and Technologies in Health (CADTH) are available [6]. In Switzerland, medical ethics guidelines regarding coercive measures in medicine exist, including recommendations for restraint use in general [7].

Restraints are used frequently in hospitals. Internationally, prevalence rates range from 0% to 100% [e.g., 8–10]. These large differences may primarily be influenced by different definitions of restraint use, the setting (e.g., intensive care units or general wards), the legal situation in a particular country, or the availability of equipment (e.g., belts for mechanical fixation) in the institution or on the ward [11–14]. In hospitals, restraints are frequently used to prevent adverse events such as falls or therapy interruption [10, 15–18]. However, the effectiveness of restraints for these purposes is increasingly being questioned. Var-

Correspondence:

Silvia Thomann
Bern University of Applied Sciences
Department of Health Professions
Applied Research and Development in Nursing
Murtenstrasse 10
CH-3008 Bern
silvia.thomann[at]bfh.ch

ious studies have reported that restraints have no effect on fall prevention or self-extubation [18–22]. In contrast, there is evidence that restraints in hospitals are associated with negative consequences for patients' physical and mental health [23–25], and also with moral distress among health professionals [18, 26]. Thus, based on the available evidence, it is not certain that in the hospital setting the benefits of restraint use exceed the harms, which is a basic ethical requirement for restraint use. Therefore, it is recommended to reduce restraint use as much as possible [6, 27]. Different measures for restraint reduction have been examined in various healthcare settings. Many studies have concluded that individual measures such as education of health professionals can be beneficial. However, national approaches might have an even greater impact [28, 29]. Local or national measurement and quality improvement initiatives are known in the nursing home and mental health care settings [3, 30–32]. However, national approaches to restraint reduction might also be relevant for hospitals. Apart from the considerable differences in prevalence rates described above, there is increasing evidence that restraint use in hospitals also depends on patient-independent factors such as routine, local habits, organisational attitudes, and hospital structures and policies [11–14, 33–35]. When such factors are recognised by hospital management and staff as being relevant to the reduction of restraint use, efforts can be made to change them.

Often, a key aspect of national programmes is the measurement and benchmarking of certain indicators of clinical performance [36]. Benchmarking can allow critical reflection upon the restraint practice within the ward or institution and the identification of potential for improvement. However, such quality measurements are only meaningful if risk-adjusted differences between hospitals (that take into account the different patient mixes) are identified, because it is these that can reveal potential for quality improvement [36].

Therefore, the aim of this study was to reanalyse existing data on restraint use in Swiss hospitals in order to assess the potential of restraint use as a national quality indicator for the hospital setting.

Materials and methods

Study design and setting

Using a cross-sectional, multicentre design, the data used for the secondary analysis were collected as part of the ANQ's (Swiss National Association for Quality Development in Hospitals and Clinics) "falls and pressure ulcers" national prevalence measurement in Switzerland [37]. The annual measurement of falls and pressure ulcers is mandatory for all hospitals in Switzerland. In addition to these two indicators, hospitals can measure restraint use on a voluntary basis. For the present study, data from hospitals measuring restraint use at the following three measurement points were included: 08 November 2016, 14 November 2017 and 13 November 2018.

Sample

The sample consisted of patients aged 18 and older who were hospitalised on one of the reference dates when the

measurement took place and who (or whose legal representative) gave informed oral consent for their participation in the overall quality measurement. The documentation of oral consent was the responsibility of the hospitals. It was recommended that consent be recorded in a central document or in the patient documentation. All medical specialties (ward types) were included, except for maternity units, emergency departments and post-anaesthesia care units. Patients who were not available on the ward during the measurement (e.g., those undergoing surgery) were excluded. We did not apply any other exclusion criteria for this secondary analysis.

Instrument and data collection

Data were collected utilising the LPZ 2.0 (Landelijke Prevalentiemeting Zorgkwaliteit) instrument (version 2016), which was developed by an international consortium led by Maastricht University in the Netherlands ([38], www.lpz-um.eu). This instrument assesses general and care indicator specific information at the institutional, ward and patient levels. In this study, we conducted a secondary data analysis using variables relating to restraint use at all three levels (see supplementary table S1 in the appendix). Restraints were defined as "interventions that may infringe [on] a person's human rights and freedom of movement, including observation, seclusion, manual restraint, mechanical restraint and rapid tranquillisation" [39]. This definition largely corresponds to that of the Swiss Academy of Medical Sciences (SAMS), from which the national language translation for the LPZ 2.0 measurement is derived [7].

The LPZ 2.0 instrument utilises standardised data collection procedures. The entire procedure (e.g., recruitment of and obtaining information from patients and preparation of data for collection, including documentation of restraint use 30 days prior to data collection) and all questions and answer options are the same across the participating LPZ consortium nations and are described in a manual. Data are collected via an online tool that guarantees completion of the questionnaire. To ensure uniform execution of the measurement across participating hospitals, data collectors were trained prior to the measurement. Utilising the train-the-trainer concept, the national coordinator trained a responsible person within each hospital (the institutional coordinator). The institutional coordinator subsequently trained the data collectors (registered nurses) on the wards. Additionally, the measurement manual containing all the relevant information was readily available (within the data entry programme) to the data collectors.

Statistical analysis

The data from the different measurement points from 2016 to 2018 were pooled. Descriptive statistics (numbers, percentages, 95% confidence intervals [CIs], medians and interquartile ranges [IQRs]) were used to describe the sample.

For the benchmarking, a multilevel modelling approach was used. This approach allows for the adjustment of "... patient-level risk factors that are outside the control of providers" [36]. In other words, the different patient mixes, and thus the risk of using more or fewer restraint, which

depend on the complexity of the patient situations in each hospital, were considered in the benchmarking. A very similar approach is used for the ANQ's "falls and pressure ulcers" national prevalence measurement [37]. The model was built as follows: restraint use was defined as the dependent variable; hospitals were considered as a random effect; and the institutional and ward level variables, as well as the general patient-level information, were considered as fixed effects (see table S1 in the appendix). The ward level could not be included as a third level in the model (i.e., a three-level model could not be built), as restraint use was assessed over a period of 30 days at the hospital level and not at the ward level. Since patient transfers from one ward to another are frequent, there was a risk of misclassification, and thus bias, when including the ward level.

Given the insufficient theoretical database available on restraint use in the hospital setting (with the partial exception of for mechanical restraint use in the intensive care setting), it was not possible to determine which of all possible fixed effects to include or exclude in a purely theory-driven manner. Since including all possible fixed effects might have led to an overadjustment, we decided on a data-driven model. For the data-driven modelling, we considered several variable selection procedures for logistic multilevel models. However, the very few software implementations available were not applicable to our problem. Therefore, we used the Akaike information criterion (AIC) [40] backwards procedure, implemented in the R package "MASS" [41]. Consequently, the hospital random effect had to be treated as a fixed effect for variable selection. In addition, the AIC procedure was employed in such a way that the hospital variable could not be unselected, as comparing between hospitals is an explicit part of this study. To reduce the number of noisy variables selected due to the large sample size and thus enhance the stability of the variable selection, we used a split-half approach, in which the data were randomly divided into two subsets and the AIC procedure was applied to each subset. For the final multilevel model, we used only the variables included in both selections. Afterwards, a generalised linear mixed model, fit by maximum likelihood (Laplace approximation), was built using the R package "lme4" [42]. To assess the relevance of the random effect, the intraclass correlation coefficient (ICC) was estimated, and a log-likelihood ratio test was performed. Afterwards, a caterpillar plot was built with all hospitals on the x-axis and their residuals and 95% CIs on the y-axis.

The ICD-10 diagnosis groups [43] "pregnancy, childbirth and the puerperium"; "congenital malformations, deformations and chromosomal abnormalities"; and "certain conditions originating in the perinatal period", as well as the answer option unknown/no diagnosis, had to be excluded, as they were present in less than 1% of the patients. The inclusion of these variables would have led to convergence problems in the variable selection. In addition, the variables "age in years" and "number of days since admission to hospital" had to be centred for similar reasons. Multicollinearity was tested using the variance inflation factor (VIF). There were no missing data, as the online data entry program only allowed the survey to be finished if all the questions were answered.

The statistical analysis was conducted using R version 4.0.1 [44] and the R packages "compareGroups" [45], "Hmisc" [46], "lme4" [42], "jtools" [47], "MASS" [41], "MuMIn" [48], "sjPlot" [49] and "tidyverse" [50].

Ethical considerations

The Ethics Committee of the Canton of Bern declared that the present study was not subject to the Swiss Human Research Act (April 2019, BASEC-Nr: Req-2019-00259). Therefore, ethical approval was not required. All patients, or their legal representatives, received written information about the measurement and gave their oral informed consent. Data were collected pseudonymously, so that no conclusions could be made regarding the individual patients. Participation was voluntary.

Results

The sample consisted of 18,938 patients who were hospitalised in 55 hospitals (table 1). The participation rate was 76.6% ($n = 18,938/24,736$) across all three years. The 30-day prevalence rate of patients with at least one restraint was 10.2% ($n = 1933/18,938$). Detailed information regarding restraint type used (e.g., mechanical, pharmacological or electronic), reason for restraint use (e.g., fall prevention or aggression) and the processes surrounding restraint use (e.g., documentation or evaluation), as well as the distribution of the sample across different hospital and ward types, are available in the appendix.

For 68.7% ($n = 13,016/18,938$) of the patients there were institutional level guidelines regarding restraint use in the respective hospital and for 34.3% ($n = 6503/18,938$) a multi-disciplinary expert committee on restraints was available. At the ward level, 66.7% ($n = 12,635/18,938$) of the patients surveyed were hospitalised in wards where regular audits were carried out to ensure compliance with the guidelines regarding restraints. Nursing staff had attended a refresher course on restraint use in 10.5% ($n = 1980/18,938$) of all patient situations.

Several factors associated with restraint use were found in the multilevel regression analysis (table 2). Patients' care dependency showed the strongest association with restraint use (odds ratio [OR] 52.65, 95% CI 41.71–66.48 for completely dependent patients compared to completely independent patients). Furthermore, a strong association between mental and behavioural disorders and restraint use was found (OR 2.22, 95% CI 1.97–2.49). No organisational factors were selected for the model.

In total, 35% of the variation in restraint use could be explained by fixed effects (selected patient characteristics; marginal $R^2 = 0.35$). The full model, including the random effect (hospital, as a cluster variable), explained 43% of the variation in restraint use (conditional $R^2 = 0.43$). Based on the ICC (0.12) and the log-likelihood ratio test (p -value < 0.000), there was relevant and significant between-hospital variability, underlining the relevance of hospital as a random effect and thus indicating the great potential for benchmarking the use of restraints across hospitals.

The risk-adjusted hospital comparison (fig. 1) showed that hospitals in Switzerland differ significantly in their use of restraints, even when adjusting for patient characteristics. Figure 1 shows how 10 hospitals differed positively (in a

clinical sense, i.e. they had lower restraint rates compared to other institutions) and 12 hospitals differed negatively (i.e., they had higher restraint rates) from the average.

Discussion

In this secondary data analysis of cross-sectional data on restraint use in Swiss hospitals, we found a prevalence of restraint use of 10.2%. We detected a strong association of restraint use with the care dependency of patients, as well as with patients with mental and behavioural disorders. Furthermore, Swiss hospitals differed significantly regarding their restraint use, even after the adjustment for risk (taking into account the different patient mixes, and thus the different risk of hospitals for using restraints). Overall, 22 out of 55 hospitals differed significantly, either positively or negatively, from the average.

The 55 participating hospitals make up about one quarter of all hospitals in Switzerland. The characteristics of the patients included are comparable to those in the mandatory fall and pressure ulcer measurements, carried out using the same methodology [37]. Consequently, it can be assumed that the sample is likely to be representative for Swiss hospitals. The prevalence of restraint use of 10.2% includes electronic measures such as sensor mats or video surveillance, whereas most other studies in the hospital setting only examined mechanical restraint with belts. Therefore, comparisons of prevalence rates are not possible. However, as reported by Thomann et al. [10], mechanical restraint with belts constitutes only 9.7% of all mechanical restraints used in Swiss hospitals.

Based on the multilevel regression analysis, a very vulnerable patient group, namely older, more care-dependent

Table 1:
Sample description.

Characteristics	Total (n = 18,938)	
	n	% (95% CI)
Institutional level		
Availability of guidelines regarding restraints (yes)	13,016	68.7 (68.1–69.4)
Availability of a multi-disciplinary expert committee on restraints (yes)	6503	34.3 (33.7–35.0)
Ward level		
Performance of regular audits to ensure compliance with the guidelines regarding restraints (yes)	12,635	66.7 (66.0–67.4)
Refresher course on restraints in the last two years for at least 80% of the ward's nursing staff (yes)	1980	10.5 (10.0–10.9)
Patient level	median	IQR
Age in years	70	24
Number of days since admission to hospital	5	9
Care dependency scale (sum score) ^a	70	15
	n	% (95% CI)
Female gender	9031	47.7 (47.0–48.4)
Surgical intervention in the two weeks prior to data collection (yes)	7667	40.5 (39.8–41.2)
Three most frequent ICD-10 diagnosis groups (multiple responses)		
Diseases of the circulatory system	10,757	56.8 (56.1–57.5)
Diseases of the musculoskeletal system and connective tissue	6829	36.1 (35.4–36.7)
Endocrine, nutritional and metabolic diseases	6432	34.0 (33.3–34.6)
Restraint use (yes)	1933	10.2 (9.8–10.6)

IQR = interquartile range; 95% CI = 95% confidence interval; ICD-10 = International Statistical Classification of Diseases and Related Health Problems 10th Revision [43].

^a Care dependency assessed using the care dependency scale (CDS) [54]. In the CDS, 15 items (e.g., eating and drinking or mobility) are rated on a Likert scale from 1 (completely dependent) to 5 (completely independent). It results in a sum score of 15–75 (higher score indicating higher care dependency), which is divided into five categories (15–24 completely dependent, 25–44 dependent to a great extent, 45–59 partially dependent, 60–69 independent to a great extent, 70–75 completely independent).

Table 2:
Multilevel logistic regression model.

Model: AIC 9025.02; marginal R ² = 0.35; conditional R ² = 0.43; ICC = 0.12	
Random effect	Variance (SD)
Hospital (intercept)	0.45 (0.67)
Fixed effects	OR (95% CI)
(Intercept)	0.02 (0.01–0.02)*
Age in years	1.01 (1.01–1.02)*
Female gender	0.71 (0.64–0.79)*
Number of days since admission to hospital	1.01 (1.01–1.01)*
Care dependency scale (CDS)	
– ≥70 completely independent	Reference
– 60–69 independent to a great extent	3.37 (2.80–4.07)*
– 45–59 partially dependent	9.74 (8.11–11.71)*
– 25–44 dependent to a great extent	27.42 (22.50–33.42)*
– ≤24 completely dependent	52.65 (41.71–66.48)*
ICD-10 diagnosis group: mental and behavioural disorders	2.22 (1.97–2.49)*
ICD-10 diagnosis group: factors influencing health status and contact with health services	1.33 (1.12–1.58)*

* Statistically significant based on the 95% CI

AIC = Akaike information criterion; ICC = intraclass correlation coefficient; SD = standard deviation; OR = odds ratio; 95% CI = 95% confidence interval; ICD-10 = International Statistical Classification of Diseases and Related Health Problems 10th Revision

patients and those with mental and behavioural disorders, seemed to be most affected by restraint use. This result is ethically highly relevant, as restraint use affects a group of patients who are often unable to defend their own rights. Therefore, it seems even more important that any use of restraints is critically analysed from both an ethical and a legal point of view. In this context, it is important to note that restraint use often violates a basic ethical principle: the expected positive health effects must exceed the harm. The positive effects of restraints in the hospital setting have not yet been proven [14, 18, 22].

Based on the risk-adjusted benchmarking, we found that restraint use differs significantly in Swiss hospitals. The caterpillar plot shows that 40% of all hospitals differ significantly, either positively or negatively, from the average. In comparison to other quality indicators, this is a very high value. For example, the same data collection method and a very similar statistical method were utilised for the ANQ's "falls and pressure ulcers" national prevalence measurement. The number of outliers in past measurements of these quality indicators has been between 0% and 8% [37]. In other words, care quality regarding falls and pressure ulcers differs only slightly between Swiss hospitals. In contrast, there are relevant differences in restraint use. Such differences indicate potential for improvement [51]. Based on the results of this study, it remains unclear how the differences can be explained, especially as none of the included structural characteristics (guidelines, expert committees, audits, refresher courses) were selected for the model. Thus, it remains unclear which quality improvement measures could be effective in reducing restraint use. Nevertheless, as mentioned above, factors that are difficult to measure, such as routine, institutional culture or attitudes, may have an influence on the results. Since restraint use is associated not only with quality of care, but also with human rights, it seems legally and ethically problematic if decision-making is based on (individual) opinions, attitudes or culture. Clearer and binding regulations and

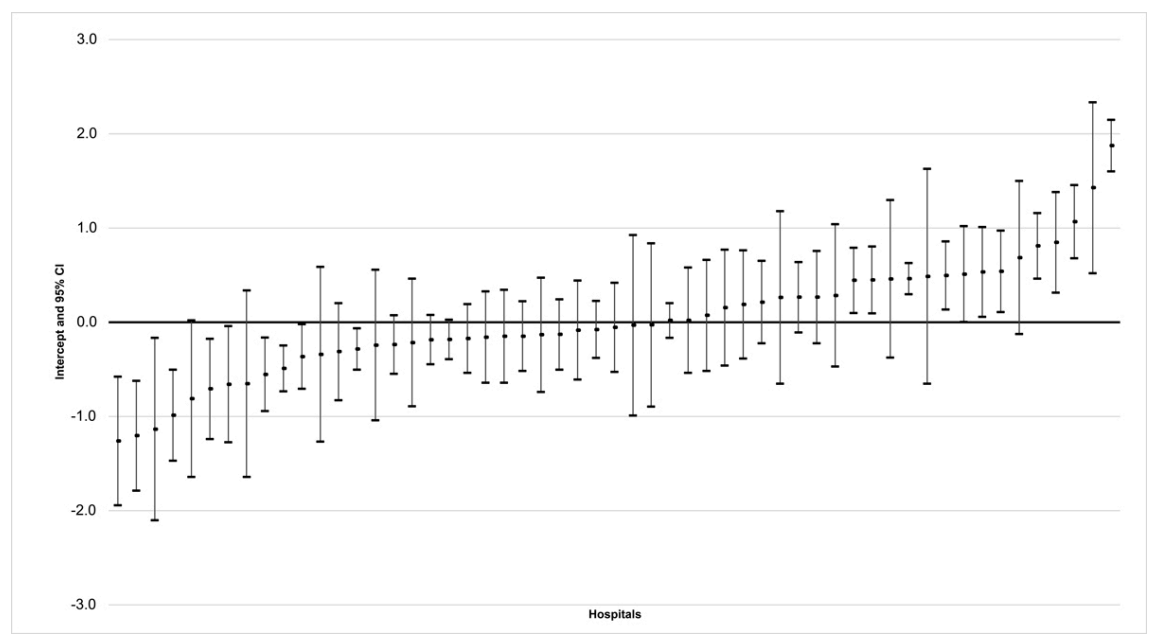
the promotion of critical scrutiny of hospitals' internal restraint practices could help to address the dilemma of legal certainty versus practical challenges (e.g., in patients with cognitive impairments) [26]. Thus, there seems to be a lot of potential for restraint use as a quality indicator for hospitals.

Restraint use is a very sensitive issue, and clear, binding legal regulations for hospitals are lacking [10]. Therefore, a national approach to quality measurement and development seems necessary. Firstly, such an approach would encourage discussion of the issue among policymakers, professional organisations and society. This would result in the establishment of the necessary structures at a macro level, which is an important element of quality development [36]. Secondly, a national approach consisting of monitoring and benchmarking would stimulate critical interprofessional discussions, both at different management levels within institutions and in direct patient care. Such interprofessional discussions across the different organisational levels of a hospital are needed to reflect on and address its institutional culture and routine, which seem to play an important role in the use of restraints [14, 52]. Thirdly, a national approach could also contribute to improving the current lack of data and evidence on restraint use in hospitals [53]. This would then enable the development and implementation of a (national) quality improvement program. Interprofessional decision-making based on evidence would then be promoted, instead of decision-making being based on personal opinions, the intuition or the institutional culture. An adequate database would also allow examination of the extent to which concepts for better restraint prevention and management from long-term care or mental health settings could be adapted for the hospital setting.

Conclusion

Although restraint use potentially violates human rights, there are no clear and binding legal regulations regarding

Figure 1: Risk-adjusted restraint use hospital comparison (residuals and 95% confidence interval [CI]).



restraint use in hospitals, despite it being well known that they are frequently used in this setting. This study highlights large risk-adjusted differences between Swiss hospitals regarding their restraint use. These differences seem questionable from a professional, ethical and legal point of view. Therefore, monitoring and benchmarking of restraint use in hospitals through a national quality indicator seems necessary. This would help ensure that restraint use is in alignment with professional values, as well as ethical and legal requirements. Additionally, it would stimulate quality improvement in this area and guarantee high-quality care in Swiss hospitals.

Limitations

One limitation is the definition of restraints used. As can be expected, not all restraint types restrict freedom and human rights to the same extent, so it would be worth examining whether different restraint types should be analysed separately. Nevertheless, even measures such as sensor mats are restrictive interventions whose effectiveness has not yet been proven [7, 22]. On the contrary, it is currently a topic of discussion as to whether the risk of undesirable events increases when such electronic measures are used without reflection, thereby causing „alarm fatigue" [22]. Furthermore, the Swiss Academy of Medical Sciences guidelines also include electronic restraints, emphasising the need for critical reflection regarding their use and, therefore, the need to measure them along with all other restraint types [7]. We were not able to comprehensively cover the diversity of restraint measures. Different restraint measures have different impacts on the affected patients, both objectively and subjectively. However, it is unclear whether, for example, the impact felt from mechanical restraints is subjectively worse than that from pharmacological interventions. Much more sophisticated research is needed to gain more insight on this matter.

A second limitation with this measurement is that a potentially very vulnerable patient group, one that is heavily affected by restraint use (older patients, more care-dependent patients and/or patients with cognitive impairments), may have been excluded, as they were not able to give informed consent. Therefore, a selection bias may exist. A third limitation is the possibility of a recall or documentation bias, as restraint use was assessed within the institution over a period of 30 days. It is known that the use of restraints is often not well documented; therefore, some undocumented restraint use may not have been assessed within this measurement [10, 18]. Consequently, the use of restraints may have been underestimated. The assessment of restraint use over a 30-day period at the hospital level also meant the ward level could not be included in the multilevel modelling (i.e., a three-level model could not be built; see the Materials and methods section). Since restraint use may differ depending on ward type, future studies should assess restraint use at the ward level, as this could provide important information regarding intra-hospital variation.

Some limitations must also be expected due to the cross-sectional design and the instrument used. The cross-sectional design favours detecting variation within the population assessed, and the detection of causal associations and/or the direction of any associations is not possible. For example, care dependency could be the reason for, but also

a consequence of restraint use. The instrument utilised included only certain organisational factors, and these were not selected for the model. In order to stimulate quality improvement, it would be worth examining which organisational factors are associated with restraint use. Due to the limited evidence available, some patient characteristics that are relevant for the risk adjustment might also be missing. A more in-depth investigation of risk factors to ensure adequate risk adjustment is necessary.

In addition, the hospital type was not considered in this analysis. However, we assumed that care dependency acted as a kind of proxy variable in this context, as the complexity of the patient cases and, consequently, the extent of (medical) care needed is relevant for the differences between hospital types. Also, due to hospital mergers, there is a risk of inadequate classification, as the hospital group's classification may not be accurate for all hospital sites within that hospital group.

Apart from these limitations, the results are likely to be generalisable for Swiss hospitals, as the large sample studied is comparable with the population assessed in the ANQ's „falls and pressure ulcers" national prevalence measurement in Swiss hospitals [37]. In addition, the data collection method is well-established in Swiss hospitals and is expected to have a positive impact on data quality.

Data availability statement

The data that support the findings of this study are available from the Swiss National Association for Quality Development in Hospitals and Clinics. However, restrictions apply to the availability of these data, which were used under licence for the current study and so are not publicly available.

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Potential competing interests

All authors have completed and submitted the International Committee of Medical Journal Editors form for disclosure of potential conflicts of interest. No potential conflict of interest was disclosed.

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Appendix

Detailed information regarding variables (table S1), hospital and ward types (table S2), restraint types (table S3), reasons for restraint use (table S4) and the processes sur-

rounding restraint use (table S5). A discussion of restraint types, reasons for restraint use and the processes surrounding restraint use can be found in Thomann et al. [10].

Table S1:

Variables.

Level and variable	Details
Institutional level	
Availability of a protocol/guidelines regarding restraints (based on national/international guidelines) within the institution	Yes/no
Availability of a multi-disciplinary expert committee on restraints within the institution	Yes/no
Ward level	
Performance of regular audits to ensure compliance with the protocol/guidelines regarding restraints	Yes/no
Assessment regarding whether at least 80% of the ward's nursing staff had attended a refresher course on restraints in the last two years	Yes/no
Patient level	
Age in years	Continuous
Sex	Female, male
Surgical intervention in the two weeks prior to data collection	Yes/no
Number of days since admission to hospital	Continuous
Medical diagnosis groups according to ICD-10 (International Statistical Classification of Diseases and Related Health Problems 10th Revision) [43]	Yes/no for each diagnosis group
Care dependency assessed using the care dependency scale (CDS) [54]. In the CDS, 15 items (e.g., eating and drinking or mobility) are rated on a Likert scale from 1 (completely dependent) to 5 (completely independent). It results in a sum score of 15–75 (higher score indicating higher care independency), which is divided into five categories (15–24 completely dependent, 25–44 dependent to a great extent, 45–59 partially dependent, 60–69 independent to a great extent, 70–75 completely independent).	Continuous or ordinal
Restraint use within the institution retrospectively over a maximum period of 30 days	Yes/no

Table S2:

Sample description – hospital and ward type.

Characteristics	Total (n = 18,938)	
	n	% (95% CI)
Hospital type^a		
Centre care hospital	8642	45.6 (44.9–46.3)
University hospital	7384	39.0 (38.3–39.7)
Primary care hospital	2537	13.4 (12.9–13.9)
Specialised hospital	375	2.0 (1.8–2.2)
Ward type		
Surgical	8576	45.3 (44.6–46.0)
Non-surgical (medical)	7154	37.8 (37.1–38.5)
Acute geriatrics	950	5.0 (4.7–5.3)
Intensive care	784	4.1 (3.9–4.4)
High dependency care	411	2.2 (2.0–2.4)
Gynaecology	409	2.2 (2.0–2.4)
Other	401	2.1 (1.9–2.3)
Short stay	147	0.8 (0.7–0.9)
Palliative care	106	0.6 (0.5–0.7)

^a hospital type (specialisation) according to the Swiss Federal Office of Public Health [55]

Table S3:

Restraint type.

Patients with restraint (n)	1933	
Restraint type (multiple responses)	n	% (95% CI)
Mechanical restraints	1125	58.2 (56.0–60.4)
<i>Type of mechanical restraint (multiple responses, only available for 2018)</i>		
<i>n participants in 2018</i>	6344	
<i>n mechanical restraints (yes) in 2018</i>	454	
Bed rails	397	87.4 (84.0–90.4)
Other mechanical restraints	85	18.7 (15.2–22.6)
Belt fixation	43	9.5 (6.9–12.5)
Tabletop / chair table	43	9.5 (6.9–12.5)
Electronic restraints	694	35.9 (33.8–38.1)
Pharmacological restraints	504	26.1 (24.1–28.1)
Other	281	14.5 (13.0–16.2)
One-to-one supervision ^a	202	10.5 (9.1–11.9)
Physical restraint (keeping someone restrained with human physical force)	67	3.5 (2.7–4.4)
Locked ward or building	57	2.9 (2.2–3.8)

95% CI = 95% confidence interval

^a This answer option was only available for 2017 and 2018 (n participants = 12,560).**Table S4:**

Reasons for restraint use.

Patients with restraint (n)	1933	
Main reason for restraint use (single response)	n	% (95% CI)
(Preventing) falls	935	48.4 (46.1–50.6)
Confusion or delirious behaviour	419	21.7 (19.9–23.6)
Other motive	190	9.8 (8.5–11.2)
Agitation	106	5.5 (4.5–6.6)
At request of the patient and/or their family	99	5.1 (4.2–6.2)
Non-compliance with treatment	68	3.5 (2.7–4.4)
(Preventing) wandering around	37	1.9 (1.4–2.6)
Unknown	18	0.9 (0.6–1.5)
(Preventing) aggressive behaviour	15	0.8 (0.4–1.3)

95% CI = 95% confidence interval

Table S5:

Processes surrounding restraint use.

Patients with restraint (n)	1933	
Process indicators (multiple responses)	n	% (95% CI)
The restraining was documented in the patient file.	1270	65.7 (63.5–67.8)
The patient and/or their legal representatives were informed about the entire process surrounding the use of restraints.	985	51.0 (48.7–53.2)
In each shift a person/nurse was appointed to monitor the patient undergoing restraint regularly, according to the defined prescription.	858	44.4 (42.2–46.6)
The use of restraints was evaluated by all persons involved (including the patient).	836	43.2 (41.0–45.5)
Primarily alternatives were used to minimise the use of restraints.	724	37.5 (35.3–39.7)
None	208	10.8 (9.4–12.2)

95% CI = 95% confidence interval